

# The Metropolis–Hastings Algorithm

## 1 Introduction

Suppose we wish to draw samples from a probability distribution  $\pi$  defined on a state space  $\mathcal{X} \subseteq \mathbb{R}^d$  for which direct sampling is impractical. Typically  $\pi$  is known only up to a normalising constant, i.e.

$$\pi(x) = \frac{1}{Z} p(x), \quad x \in \mathcal{X}, \quad Z = \int_{\mathcal{X}} p(x) dx,$$

where the non-negative function  $p$  can be evaluated point-wise, but the integral  $Z$  is intractable.

The *Metropolis–Hastings* (MH) algorithm constructs a Markov chain whose stationary distribution is  $\pi$ , yielding asymptotically exact samples via simulation.

## 2 Algorithm

Let  $q(y|x)$  be a *proposal (transition) kernel*, i.e. a conditional probability density (or mass) function from which we can sample efficiently. Common choices include symmetric random walks and independence proposals.

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**Algorithm 1** Metropolis–Hastings

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**Require:** unnormalised density  $p$ , proposal kernel  $q$ , initial state  $x_1$

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1: for  $t = 1, 2, \dots$  do
2:   Sample  $y \sim q(\cdot | x_t)$  ▷ proposal step
3:   Compute the acceptance probability

$$a = \min\left(1, \frac{p(y) q(x_t | y)}{p(x_t) q(y | x_t)}\right).$$

4:   Draw  $u \sim \text{Uniform}(0, 1)$ 
5:   if  $u < a$  then
6:      $x_{t+1} \leftarrow y$  ▷ accept
7:   else
8:      $x_{t+1} \leftarrow x_t$  ▷ reject
9:   end if
10: end for
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### 2.1 Examples of proposals

**Gaussian random walk.**  $q(y|x) = \mathcal{N}(x, \sigma^2 I)$ , where  $\sigma$  controls the step size. This proposal is symmetric, i.e.  $q(y|x) = q(x|y)$ . In this case, the acceptance probability is simply

$$a = \min\left(1, \frac{p(y)}{p(x_t)}\right).$$

This is the original Metropolis algorithm, which is a special case of MH.

**Independence proposal.**  $q(y|x) = q(y)$ , independent of  $x$ . The acceptance probability becomes

$$a = \min\left(1, \frac{p(y)q(x_t)}{p(x_t)q(y)}\right).$$

Independence proposals can be efficient if  $q$  is a good approximation to  $\pi$ , but may lead to low acceptance rates if  $q$  is poorly chosen. Independence proposals are also not symmetric.

### 3 Convergence

#### 3.1 Detailed balance

Algorithm 1 defines a Markov chain with transition kernel  $K(x, dy)$  given by

$$K(x, dy) = q(y|x) A(x, y) dy + r(x) \delta_x(dy),$$

where

$$A(x, y) = \min\left(1, \frac{p(y)q(x|y)}{p(x)q(y|x)}\right), \quad r(x) = 1 - \int_{\mathcal{X}} q(z|x) A(x, z) dz,$$

and  $\delta_x$  denotes the Dirac measure at  $x$ . The first term describes proposed moves that are accepted, while the second term corresponds to rejected moves, ensuring that  $K(x, \cdot)$  is a probability measure.

**Proposition 1** (Detailed balance). *The MH transition kernel  $K$  satisfies the detailed balance condition with respect to  $\pi$ , i.e.*

$$\pi(dx) K(x, dy) = \pi(dy) K(y, dx) \quad \text{on } \mathcal{X} \times \mathcal{X}.$$

*Proof.* We verify the equality separately on the off-diagonal set  $\{(x, y) : x \neq y\}$  and on the diagonal.

*Off-diagonal part.* For  $x \neq y$ , we have

$$K(x, dy) = q(y|x) A(x, y) dy,$$

so that

$$\pi(dx) K(x, dy) = \pi(x) q(y|x) A(x, y) dx dy.$$

Using the definition of  $A$  and the identity  $\min(1, r) = r \min(1, 1/r)$ , we obtain

$$\begin{aligned} \pi(x) q(y|x) A(x, y) &= \pi(x) q(y|x) \frac{p(y)q(x|y)}{p(x)q(y|x)} A(y, x) \\ &= \frac{p(y)}{Z} q(x|y) A(y, x) \\ &= \pi(y) q(x|y) A(y, x), \end{aligned}$$

which shows that

$$\pi(dx) K(x, dy) = \pi(dy) K(y, dx) \quad \text{for } x \neq y.$$

*Diagonal part.* On the diagonal, the kernel is given by the rejection mass

$$K(x, dy) = r(x) \delta_x(dy),$$

so that

$$\pi(dx) K(x, dy) = \pi(dx) r(x) \delta_x(dy).$$

By symmetry, we also have

$$\pi(dy) K(y, dx) = \pi(dy) r(y) \delta_y(dx).$$

Both measures are supported on the diagonal  $\{(x, x) : x \in \mathcal{X}\}$  and, for any measurable set  $B \subset \mathcal{X} \times \mathcal{X}$ ,

$$\int \mathbf{1}_B(x, y) \pi(dx) r(x) \delta_x(dy) = \int \mathbf{1}_B(x, x) \pi(dx) r(x),$$

which is identical to the corresponding expression with  $x$  replaced by  $y$ . Hence the two measures coincide on the diagonal.

Combining the off-diagonal and diagonal parts yields

$$\pi(dx) K(x, dy) = \pi(dy) K(y, dx),$$

completing the proof. □

### 3.2 Stationarity and Convergence

The detailed balance condition implies that  $\pi$  is *invariant* for the Markov kernel  $K$ , i.e.

$$\int_{\mathcal{X}} \pi(dx) K(x, dy) = \pi(dy).$$

Equivalently, if  $X_n \sim \pi$ , then  $X_{n+1} \sim \pi$ .

To obtain stronger properties such as uniqueness and convergence, additional assumptions on the chain are required. In general state spaces, the appropriate notion of irreducibility is  $\pi$ -irreducibility, and recurrence must be understood in the sense of Harris.

Assume that the chain is  $\pi$ -irreducible and Harris recurrent. Then  $\pi$  is the *unique* invariant probability measure.

A sufficient condition for  $\pi$ -irreducibility is that the proposal kernel satisfies

$$q(y | x) > 0 \quad \text{for all } x, y \in \text{supp}(\pi),$$

so that every set of positive  $\pi$ -measure can be reached with positive probability. More generally, it suffices that for any measurable set  $A \subset \mathcal{X}$  with  $\pi(A) > 0$  and any  $x \in \mathcal{X}$ , there exists  $n \geq 1$  such that  $K^n(x, A) > 0$ .

If, in addition, the chain is aperiodic, then for any initial distribution, the law of  $X_n$  converges to  $\pi$  in total variation.

Aperiodicity typically follows from the presence of self-transitions. In the Metropolis–Hastings algorithm,

$$K(x, \{x\}) = 1 - \int_{\mathcal{X}} q(y | x) A(x, y) dy,$$

which is strictly positive whenever proposals are rejected with positive probability. In that case, the chain cannot be trapped in deterministic cycles.

Under the above conditions, the ergodic theorem holds: for any measurable function  $h : \mathcal{X} \rightarrow \mathbb{R}$  with  $h \in L^1(\pi)$ ,

$$\frac{1}{T} \sum_{t=1}^T h(X_t) \xrightarrow{T \rightarrow \infty} \int_{\mathcal{X}} h(x) \pi(dx) \quad \text{almost surely.}$$

Thus, empirical averages along the MH trajectory provide consistent estimators of expectations under  $\pi$ .